# Large deviations for weighted empirical measures from importance sampling

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How can large deviation results on the empirical measure level be useful for theoretical analysis of efficiency of importance sampling algorithms?

Performance of simulation algorithms

### Problem setting

Let X be a random variable, distribution F, taking values in some space  $\mathcal{X}$ . Consider the task of computing  $\Phi(F)$  for some functional Ф:

- **Expectation**:  $\Phi_h(F) = \int h dF =: F(h)$ , for some  $h: \mathcal{X} \mapsto \mathcal{R}$ ,
- Quantile:  $\Phi_{q}(F) = F^{-1}(q) = \inf\{x : F((x, \infty)) < q\},\$  $q \in (0,1),$
- L-statistic:  $\Phi(F) = \int_0^1 \phi(q) F^{-1}(q) dq$ .

Performance of simulation algorithms

When explicit computation is impossible, turn to simulation.

### Introduction

#### Standard Monte Carlo and importance sampling

■ Standard Monte Carlo: Sample  $X_1, ..., X_n$  i.i.d. from F and construct the empirical measure

$$\mathbb{F}_n = \frac{1}{n} \sum_{i=1}^n \delta_{X_i}.$$

■ Importance sampling: Sample  $X_1,...,X_n$  i.i.d. from the sampling distribution  $G, F \ll G$ . Construct the weighted empirical measure,

$$\mathbb{G}_n^w = \frac{1}{n} \sum_{i=1}^n w(X_i) \delta_{X_i}, \ w := \frac{dF}{dG}.$$

**Corresponding estimates are**  $\Phi(\mathbb{F}_n)$  **and**  $\Phi(\mathbb{G}_n^w)$ 

### Performance analysis

#### Introduction

- Performance of importance sampling determined by the choice of sampling distribution G. Typically evaluated in terms of  $Var(\Phi(\mathbb{G}_n^w))$
- Estimation of large deviation probabilities a well studied problem. Large deviation results have been used extensively for studying the rate of the decay of the variance and designing efficient algorithms

### Performance analysis

#### Main idea

- Rather than sample path large deviations, use large deviation results for the empirical measures.
- Replace variance by the rate function of a large deviation principle as the number of particles increases (large sample limit).

### Performance of Monte Carlo

#### Estimation of an expectation

Let  $A_{\epsilon}(F(h)) = \{x : |x - F(h)| > \epsilon F(h)\}$ . For  $\Phi(\mathbb{F}_n) = \mathbb{F}_n(h)$ , Cramér's theorem provides, for n sufficiently large, the approximation

$$P(\mathbb{F}_n(h) \in A_{\epsilon}(F(h))) \approx \exp\{-n \inf_{x \in A_{\epsilon}(F(h))} \Lambda^*(x)\}.$$

Number of particles roughly needed for an upper bound  $\delta$  on the probability:

$$n \approx \frac{1}{\inf_{x \in A_{\epsilon}(F(h))} \Lambda^{*}(x)} (-\log \delta).$$

### Performance of Monte Carlo

#### Estimation of a general functional

- For general  $\Phi$  the idea is to consider the probability of  $\mathbb{F}_n$  being close to F. Sanov's theorem provides an LDP for the empirical measures  $\mathbb{F}_n$  of Monte Carlo.
- $A_{\epsilon} \subset \mathcal{M}_1$  a set that relates to the accuracy of  $\Phi(\mathbb{F}_n)$ . By Sanov's theorem,

$$\mathbf{P}(\mathbb{F}_n \in A_{\epsilon}) \approx \exp\{-n\inf_{G \in \overline{A}_{\epsilon}} \mathcal{H}(G \mid F)\}.$$

■ Example:  $A_{\epsilon} = \{G \in \mathcal{M}_1 : | \Phi(G) - \Phi(F) | > \epsilon \Phi(F) \}.$ 

### Performance of Monte Carlo

#### Estimation of a probability

- Let *h* be the indicator of some set *A*, with p := F(A).
- With  $A_{\epsilon}$  the ball of radius  $\epsilon p$  centered at p,

$$\inf_{G\in A_{\epsilon}}\mathcal{H}(G|F)\sim lpha(\epsilon)p$$
 as  $p o 0$ .

Consistent with analysis of variance of estimator.

#### A possible approach

- Suppose  $\mathbb{G}_n^w$  satisfies an LDP. Let  $A_{\epsilon} \subset \mathcal{M}$  be some set that relates to the accuracy of the estimate  $\Phi(\mathbb{G}_n^w)$ .
- The LDP implies, for sufficiently large n,

$$\mathsf{P}(\mathbb{G}_n^w \in A_{\epsilon}) \approx \exp\{-n\inf_{\nu \in \overline{A}_{\epsilon}} I^w(\nu)\}.$$

■ With  $\delta$  the desired upper bound for the probability,

$$n pprox rac{1}{\inf_{
u \in \overline{A}_{\epsilon}} I^{w}(
u)} (-\log \delta).$$



#### A possible approach

- Similarly, the probability of  $\mathbb{G}_n^w$  having some undesireable shape  $\nu \in \mathcal{M}$  can be studied using  $I^w(\nu)$ .
- For  $\mathbb{G}_n^w(h)$  Cramér's theorem is applicable.
- Sanov's theorem not applicable for the weighted empirical measures  $\mathbb{G}_n^w$ .
- Need an LDP for  $\mathbb{G}_n^w$  in order to quantify the notion of the weighted empirical measures being close to F.

### Large deviations for importance sampling

#### Framework

- Suffices to have the weighted empirical measures  $\mathbb{G}_n^w$  close to F in the region that largely determines  $\Phi(F)$ .
- Let f be an F-integrable function characterizing the importance of different regions of the space  $\mathcal{X}$  importance function. Want

$$\mathbb{G}_n^{wf} = \frac{1}{n} \sum_{i=1}^n w(X_i) f(X_i) \delta_{X_i},$$

to be close to  $F^f$ , where  $F^f$  is defined as

$$F^{f}(g) = \int g(x)f(x)dF(x),$$

for each bounded, measurable function g.

### Large deviations for importance sampling

### Laplace principle

- $\blacksquare \ \Psi : \Gamma \mapsto \mathcal{M} \ \text{the mapping} \ G \mapsto G^{wf}.$
- $I(\nu) = \inf \{ \mathcal{H}(Q \mid G) : \Psi(Q) = \nu, Q \in \Gamma \}.$

**Theorem** Let F, G and f be given as above, with  $F \ll G$  on the support of f. Suppose that  $\int e^{wf} dG$ ,  $\int e^{w^2 f^2} dG < \infty$ . Then, for any bounded, continuous  $h : \mathcal{M} \mapsto \mathcal{R}$ ,

$$\lim_{n} \frac{1}{n} \log \mathbb{E}[e^{-nh(\mathbb{G}_n^{wf})}] = -\inf_{\nu \in \mathcal{M}} \{h(\nu) + I(\nu)\}.$$



Applications

### Performance of importance sampling

Performance of simulation algorithms

#### Performance criteria

- The choice of  $A_{\epsilon}$ , or  $\nu$ , reflects your criteria for good performance.
- Suppose that  $A \subset \mathcal{X}$  is the region of interest,  $\delta \propto F(A)$ . A possible choice is

$$A_{\epsilon} = \{ \nu \in \mathcal{M} : |\frac{d\nu}{dF}(x) - 1| \ge \epsilon \text{ for } x \in \text{ some } C \subset A, F(C) \ge \delta \}$$

Estimation of a probability

- Let  $X_1, X_2, ...$  be i.i.d. F and consider the random walk  $S_n = \sum_{i=1}^n X_i$ . Want to estimate  $p_n = \mathbf{P}(S_n/n \ge a)$  for  $\mathbb{E}[X_1] < a$ .
- Use constant exponential tilting, parameter  $\theta$ , and the previous choice of  $A_{\epsilon}$ ;  $A = \{S_n/n \geq a\}$ .
- Value of the rate function related to  $\mathbb{E}[w^{-1}(X)1\{X \in A\}]$ . Obtain a lower bound which is maximized for  $\theta$  such that

$$\kappa'(\theta) = a$$
.



### Performance of Monte Carlo methods

Rare event limit

 For both Monte Carlo and importance sampling, study the rare event limit,

$$\liminf_{p\to 0}\inf_{\nu\in A_{\epsilon}}I(\nu).$$

If the above is zero, use the asymptotic rate  $\gamma(p)$  as a measure of efficiency:

$$\inf_{
u \in A_{\epsilon}} I(
u) \sim \gamma(p) \text{ as } p \to 0.$$

### Applications for performance analysis

#### Possible ways to use the large deviation result:

- Comparison of Monte Carlo and importance sampling in terms of the rate functions of large deviation principles.
- Larger rate suggests improved performance.
- The large deviation heuristics of the most likely way for an event to occur can possibly help in designing algorithms that meet the specified performance criteria

### Summary

- Propose a way to use the rate function of large deviation results to quantify the performance of importance sampling algorithms.
- Derive a Laplace principle for the weighted empirical measures of importance sampling as the number of particles increases.

## Large deviations for importance sampling Idea of proof

- Relies on the weak convergence approach to large deviations<sup>1</sup>.
- Identify  $W_n = -\frac{1}{n} \log \mathbb{E}[\exp\{(-nh(\mathbb{G}_n^{wf})\}]$  as the total cost of a stochastic control problem and derive a representation formula.
- The Laplace principle upper bound

$$\limsup_{n} \frac{1}{n} \log \mathbb{E}[e^{-nh(\mathbb{G}_{n}^{wf})}] \leq -\inf_{\nu \in \mathcal{M}} \{h(\nu) + I(\nu)\},$$

requires the most work compared to the case of ordinary empirical measures (Sanov's theorem).



<sup>&</sup>lt;sup>1</sup>Dupuis and Ellis (1997)

Estimation of a probability, cont'd.

- With the described choice of  $A_{\epsilon}$ , minimizing the rate I corresponds to minimizing G(C) for  $C \subset A$ ,  $F(C) \geq \delta$  such that the condition on the Radon-Nikodym derivative is fulfilled.
- lacksquare Possible to explicitly characterize the optimal  $\tilde{C}$  in terms of the weight function.
- Optimal rate is the obtained by finding the G, within some prescribed family, which maximizes

$$G(\tilde{C}) = \mathbb{E}[w^{-1}(X)1\{X \in A\}].$$

